

Association of Alendronate and Risk of Cardiovascular Events in Patients With Hip Fracture

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ABSTRACT

The risk of cardiovascular events (CVEs) with alendronate use in real-world hip fracture patients is unknown. This study aimed to investigate the risk of CVE with and without use of alendronate in patients with hip fracture. We conducted a retrospective cohort study using a population-wide database managed by the Hong Kong Hospital Authority. Patients newly diagnosed with hip fracture from 2005 through 2013 were followed until November 6, 2016. Alendronate and other antiosteoporosis medications use during the study period were examined. We matched treated and nontreated patients based on time-dependent propensity score. The risks of cardiovascular mortality, myocardial infarction, and stroke between treatment groups were evaluated using conditional Cox regression stratified by match pairs. To examine the associations over time, outcomes were assessed at 1 year, 3 years, 5 years, and 10 years. Among 34,991 patients with newly diagnosed hip fracture, 4602 (13.2%) received antiosteoporosis treatment during follow-up. Physical functioning or survival prospect was not significantly different between treated and nontreated patients. A total of 4594 treated patients were matched with 13.568 nontreated patients. Results of Cox regression analysis revealed that alendronate was associated with a significantly lower risk of 1-year cardiovascular mortality (HR 0.33; 95% CI, 0.17 to 0.65) and incident myocardial infarction (HR 0.55; 95% Cl, 0.34 to 0.89), whereas marginally significant reduction in risk of stroke was observed at 5 years and 10 years (HR at 5 years: 0.82; 95% CI, 0.67 to 1.00; p = 0.049; HR at 10 years: 0.83; 95% CI, 0.69 to 1.01; p = 0.065). The strength of the association declined over time but remained significant. Similar results were observed when all nitrogen-containing bisphosphonates (N-BPs) were analyzed together. These findings were robust in multiple sensitivity analyses. Additional studies in other population samples and randomized clinical trials may be warranted to further understand the relationship between use of various antiosteoporosis medication and risk of CVE in patients with hip fracture. © 2018 American Society for Bone and Mineral Research.

KEY WORDS: DISEASE AND DISORDERS OF/RELATED TO BONE; EPIDEMIOLOGY; GENERAL POPULATION STUDIES; THERAPEUTICS; ANTIRESORPTIVES

Introduction

ip fracture is a common condition that leads to great morbidity and mortality in the elderly population. One of the consequences of hip fracture is an increased risk of cardiovascular events (CVEs) such as myocardial infarction (MI),⁽¹⁾ stroke,⁽²⁾ and cardiovascular mortality.^(3,4) Thus, there is a clinical need to be aware of this increased risk of CVE among patients who sustain a hip fracture, and to intervene to reduce these life-threatening outcomes. Nonetheless there are no clinical recommendations that address this issue.

Nitrogen-containing bisphosphonates (N-BPs) are the recommended treatment for the secondary prevention of fractures in persons who have sustained a fragility fracture.⁽⁵⁾ However, N-BPs are underused worldwide because of patients' concerns about potential side effects.⁽⁶⁾ Emerging evidence has suggested that N-BPs are potential cardiac-protecting agents.^(7,8) A longitudinal cohort study in women showed an association of

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N-BPs and decreased prevalence of cardiovascular calcification in older subjects.⁽⁷⁾ Another randomized clinical trial demonstrated that treatment with alendronate inhibited the progression of aortic calcification after kidney transplant, compared with no treatment with bisphosphonates.⁽⁸⁾ Animal studies found that farnesyl pyrophosphate synthase (FPPS), the molecular target of N-BPs, is involved in the pathogenesis of cardiac hypertrophy. Cardiac-specific overexpression or inhibition (using alendronate) of FPPS in mice has been shown to result in⁽⁹⁾ or attenuate⁽¹⁰⁾ cardiac hypertrophy, respectively.

Although the risk of all-cause mortality was reduced by 10% to 60% in patients treated with N-BPs after hip fracture,⁽¹¹⁻¹³⁾ the risk of CVE was inconclusive. A previous randomized controlled trial (RCT) and cohort study showed a trend toward reduction of cardiovascular mortality in patients treated with zoledronate⁽⁴⁾ or risedronate⁽¹⁴⁾ following a hip fracture. A recent meta-analysis of RCTs reported a lower risk of cardiovascular mortality in the use of bisphosphonates, although not statistically significant.⁽¹¹⁾ In view of these findings, further studies are needed to evaluate the role of antiosteoporosis medication in CVE.⁽¹⁵⁾ Because RCT data are limited with regard to CVE outcomes and participants in clinical trials are rarely representative of the actual patient population receiving medications, large observational studies that include methods to minimize confounding by indication may add important data complementing the randomized trials.

This population-based cohort study used data from a large territory-wide healthcare database to determine the risk of CVE in patients with hip fracture, with and without use of alendronate.

Materials and Methods

Methods

The study protocol was approved by the institutional review boards of the University of Hong Kong and Hong Kong HA.

Data source

Data was collected from the Clinical Data Analysis and Reporting System (CDARS), an electronic medical database managed by the Hong Kong HA. HA is a public healthcare provider that manages 42 hospitals and institutions, and 120 outpatient clinics, serving >80% of hospital admissions in Hong Kong. CDARS is a centralized database developed for the purposes of research and audit. It includes records of demographics, admission, prescription, diagnosis, procedures, laboratory tests, and deaths. All records are anonymized. The database has been widely used in conducting high-quality population-based studies^(16,17) and is specifically validated for study of the effects of medication on bone fractures.⁽¹⁸⁾ More information about Hong Kong HA is provided in the Supporting Methods.

Study cohort

This was a retrospective cohort study. We identified patients aged \geq 50 years who were admitted via an emergency room between January 1, 2005 and December 31, 2013 with a new diagnosis of hip fracture (International Classification of Diseases, Ninth Revision [ICD-9], 820.XX). Patients who survived and were discharged were included in the study cohort. To reduce selection bias and/or competing risk of death, we excluded

patients who fulfilled any of the following criteria: (i) previous exposure to antiosteoporosis medications 2 years preceding the index date; (ii) length of stay (LOS) in hospital >60 days (Supporting Methods)—patients with a longer length of stay may be less healthy and unable to take antiosteoporosis medications, so inclusion of these patients could lead to substantial selection bias—and (iii) history of cancer where antiresorptive agents are often prescribed.

Exposure and outcomes

The primary drug of interest was alendronate, which is the firstline therapy for osteoporosis, following hospital discharge postfracture. Patients were classified as "alendronate-treated" if they had at least one prescription record of alendronate before the end of the study (November 6, 2016). Bisphosphonates can accumulate in the skeleton⁽¹⁹⁾ and studies have reported a residual effect of alendronate after treatment withdrawal for up to 7 years.^(20,21) Therefore, once being treated, the patients were considered exposed to the drug until the end of follow-up. In a secondary analysis, we aimed to determine whether the association was also observed for all N-BPs as a single group (including alendronate, ibandronate, risedronate, and zoledronate), and for two antiosteoporosis medications with different mechanism-of-actions commonly prescribed in Hong Kong (>1% usage among hip fracture patients), namely strontium ranelate and salmon calcitonin (salcatonin).

Primary outcomes of interest were cardiovascular mortality, incident MI, and stroke during the follow-up period. Our previous studies validated the coding of MI and stroke in CDARS with a positive predictive value (PPV) of 85.4% and 91.1%, respectively.⁽¹⁶⁾ In the analysis of incident CVE, patients with outcomes of interest at baseline were excluded. All outcomes were defined by ICD-10 and ICD-9 and are shown in the (Supporting Tables 1 and 2).

Statistical analysis

The statistical analyses were conducted by two co-authors (CWS and AYSW) independently and cross-checked for quality assurance. Continuous variables are presented as mean \pm standard deviation (SD) and categorical variables as percentages. Incident rates per 10,000 person-years and the 95% confidence intervals (CIs) for CVE were estimated using a Poisson distribution.

Time-to-event analysis was used to evaluate the association of antiosteoporosis medication with outcomes. Because there may have been a delay in prescribing alendronate, immortal time bias was possible. Such bias, which would favor the treatment group, has been discussed elsewhere.^(22,23) To address this issue, a time-dependent propensity score matching was used,^(24,25) which matches a patient treated at time t (defined as number of days from the date of discharge to first treatment) to another patient who had not received treatment yet at time t based on the propensity score (PS) at time t. The matched pair was followed from time t until the occurrence of an event, switch to another antiosteoporosis medication, death, or study end (November 6, 2016), whichever occurred first. Using this approach, treated and nontreated groups were followed at the same starting point (time t), which has been shown to be a superior approach to control immortal time bias.⁽²⁶⁾ As exposure to treatment is time-dependent, propensity scores at different time points were estimated using Cox regression with timedependent covariates. Details are provided in Supporting

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		Any ost medicatio	eoporosis on-exposed	Alendrona	ite-exposed		Any osteo medication	oporosis 1-exposed	Alendrona	te-exposed
Characteristic	Non-exposed		Absolute standardized difference (%) ^a		Absolute standardized difference (%) ^a	Non-exposed	9 . 9	Absolute standardized ifference (%) ^a		Absolute standardized difference (%) ⁶
Subjects, <i>n</i> Males, <i>n</i> (%)	34,940 10,648 (30.5)	4602 915 (19.9)	24.6	3086 628 (20.3)	23.4	13,568 2859 (21.1) 700 (0.7)	4594 915 (19.9)	2.9	3081 628 (20.4)	0.3
Age (years), mean (SU) Age, <i>n</i> (%)	81.6 (9.4)	80.7 (8.7)	10.3	80.0 (8.8)	18.0	(1.6) 6.61	80.7 (8.7)	8.0	80.0 (8.8)	0.1
50–69 years	3594 (10.3)	502 (10.9) 3480 (75.6)		384 (12.4) 2355 (76.3)		1865 (13.7) 0807 (73.3)	502 (10.9)		384 (12.5) 2350 (76 3)	
90+ years	6710 (19.2)	(0.07) 0040 (13.5)		347 (11.2)		(0.27) 1896 1896 (14.0)	(13.5) 620 (13.5)		347 (11.3)	
Year of index date, <i>n</i> (%)			83.5		86.7			63.2		62.1
2005	3309 (9.5)	45 (1.0)		34 (1.1)		576 (4.2)	45 (1.0)		34 (1.1)	
2006	3613 (10.3)	120 (2.6)		74 (2.4)		855 (6.3)	120 (2.6)		74 (2.4)	
2007	3704 (10.6) 3065 (11 4)	206 (4.5) 334 (7.3)		122 (4.0) 165 (5 2)		1113 (8.2) 1528 (11 2)	206 (4.5) 334 (7 3)		122 (4.0) 165 (5 4)	
2000	3967 (11 4)	(c.) +cc (f1 (133)		(C.C) CO1		(C.11) 07C1 (0.11) 07C1	(c: /) +cc (133)		360 (11 7)	
2000	3870 (11.1)	(0.21) 110		552 (17.9)		2149 (15.8)	785 (17.1)		550 (17.9)	
2011	4082 (11.7)	597 (13.0)		466 (15.1)		1786 (13.2)	597 (13.0)		466 (15.1)	
2012	4011 (11.5)	548 (11.9)		382 (12.4)		1680 (12.4)	547 (11.9)		382 (12.4)	
2013	3958 (11.3)	608 (13.2)		456 (14.8)		1670 (12.3)	607 (13.2)		455 (14.8)	
2014	431 (1.2)	414 (9.0)		278 (9.0)		209 (1.5)	414 (9.0)		278 (9.0)	
2015	0 (0.0)	201 (4.4)		119 (3.9)		0 (0.0)	201 (4.4)		119 (3.9)	
2016	0 (0.0)	130 (2.8)		78 (2.5)		0 (0.0)	128 (2.8)		76 (2.5)	
Medical history, <i>n</i> (%)										
Coronary heart disease	4174 (11.9)	503 (10.9)	3.2	308 (10.0)	6.3	1362 (10.0)	502 (10.9)	2.9	308 (10.0)	0.2
Congestive heart failure	3569 (10.2)	412 (9.0)	4.3	228 (7.4)	10	1017 (7.5)	409 (8.9)	5.1	227 (7.4)	< 0.001
Cerebrovascular disease	5586 (16.0)	649 (14.1)	5.3	420 (13.6)	6.7	1799 (13.3)	649 (14.1)	2.5	420 (13.6)	2
Hypertensive disease	14,059 (40.2)	2045 (44.4)	8.5	1355 (43.9)	7.4	5570 (41.1)	2041 (44.4)	6.8	1353 (43.9)	4.6
Arrhythmia and conduction	4127 (11.8)	529 (11.5)	-	322 (10.4)	4.4	1391 (10.3)	528 (11.5)	4	321 (10.4)	0.5
ursoruers Arterial disease	1535 (44)	185 (40)	19	120 (3 9)	25	474 (35)	184 (4 U)	7 6	119 (3 9)	16
Chronic obstructive nulmonary	3308 (9.5)	464 (101)	5 1 C	(2.2) 021	2 0	(0.0) - (1)	462 (101)	, v. v.	798 (9.7)	2:- C
disease			- i		2			2		1
Diabetes	7273 (20.8)	1069 (23.2)	5.8	721 (23.4)	6.1	2956 (21.8)	1067 (23.2)	3.4	720 (23.4)	2.3
Chronic liver disease	250 (0.7)	32 (0.7)	0.2	20 (0.6)	0.8	71 (0.5)	32 (0.7)	2.2	20 (0.6)	<0.1
Renal failure	1719 (4.9)	151 (3.3)	8.3	86 (2.8)	11.1	346 (2.6)	149 (3.2)	4.1	85 (2.8)	1.4
Connective tissue disease	232 (0.7)	95 (2.1)	12.1	68 (2.2)	13	167 (1.2)	89 (1.9)	5.7	64 (2.1)	7
Dementia	2864 (8.2)	218 (4.7)	14.1	124 (4.0)	17.5	519 (3.8)	216 (4.7)	4.3	123 (4.0)	1.7
										continued

Table 1. Baseline Characteristics of the Study Population Before and After Propensity Score Matching

Table 1. (Continued)										
		₽.	rematched coho	ort				Matched coho	ч	
		Any ost medicatio	eoporosis on-exposed	Alendron	ate-exposed		Any ost medicatic	eoporosis n-exposed	Alendron	ite-exposed
	•		Absolute standardized		Absolute standardized	•		Absolute standardized		Absolute standardized
Characteristic	Non-exposed		difference (%) ^a		difference (%) ^a	Non-exposed		difference (%) ^a	_	difference (%) ^a
Paget's disease of bone	(0.0) 6	0 (0.0)	2.3	0 (0.0)	2.3	0 (0.0)	0 (0.0)	I	0 (0.0)	I
Osteoporosis	1628 (4.7)	623 (13.5)	31.2	386 (12.5)	28.3	1036 (7.6)	618 (13.5)	19	383 (12.4)	14.3
Fall	34,007 (97.3)	4519 (98.2)	5.9	3025 (98.0)	4.6	13,288 (97.9)	4511 (98.2)	1.9	3020 (98.0)	< 0.001
Other major fractures	3477 (10.0)	661 (14.4)	13.5	414 (13.4)	10.8	1703 (12.6)	661 (14.4)	5.4	414 (13.4)	1.6
Obesity	58 (0.2)	11 (0.2)	1.6	9 (0.3)	2.6	21 (0.2)	11 (0.2)	1.9	9 (0.3)	2.4
Hyperlipidemia	2716 (7.8)	480 (10.4)	9.2	336 (10.9)	10.7	1256 (9.3)	479 (10.4)	3.9	336 (10.9)	<0.1
Thyroid disorders	799 (2.3)	127 (2.8)	m	72 (2.3)	0.3	318 (2.3)	127 (2.8)	2.7	72 (2.3)	0.1
Number of outpatient visits in the	8.0 (10.2)	13.9 (12.7)	51.3	13.9 (12.6)	51.5	8.1 (10.8)	13.9 (12.7)	49.8	13.9 (12.6)	49.8
past 1 year, mean (SD)										
Prescription in past 180 days, <i>n</i> (%) Diagona) 1377 (2 8)	(0 () (21	۲ ٦	(C C) 29	0 2	378 (7 8)	(0 () (2)	2	((() 2)	ر د د
Loop diuratics	5373 (154)	(6.2) 201 646 (14 0)	- 6	399 (129)	0. r	1792 (13.2)	(6.2) 201 643 (14 0)	0.0	397 (129)	4.C 5.C
Other diritatics	3118 (8 0)	395 (86)	5.5 C 1	758 (8 4)	, r	1316 (97)	303 (86)	C:	257 (8 3)	2"- 7 P
Anti-arrhythmics class I and II	718 (2.1)	75 (1.6)	3.2	41 (1.3)	2 5.6	263 (1.9)	75 (1.6)	2.3	41 (1.3)	j. ru
Beta blockers	8676 (24.8)	1176 (25.6)	1.7	770 (25.0)	0.3	3674 (27.1)	1170 (25.5)	3.7	765 (24.8)	5.5
Angiotensin receptor blocker/	9395 (26.9)	1372 (29.8)	6.5	912 (29.6)	5.9	3925 (28.9)	1370 (29.8)	5	910 (29.5)	< 0.001
angiotensin converting										
enzyme inhibitor/renin										
inhibitor										
Nitrates	4216 (12.1)	505 (11.0)	3.4	315 (10.2)	5.9	1558 (11.5)	504 (11.0)	1.6	315 (10.2)	4.1
Calcium channel blockers	15,302 (43.8)	2078 (45.2)	2.7	1385 (44.9)	2.2	6112 (45.0)	2073 (45.1)	0.2	1382 (44.9)	2
Peripheral vasodilators	742 (2.1)	57 (1.2)	6.9	41 (1.3)	6.1	158 (1.2)	56 (1.2)	0.5	40 (1.3)	2.5
Anticoagulants	2657 (7.6)	310 (6.7)	3.4	224 (7.3)	1.3	1176 (8.7)	307 (6.7)	7.5	222 (7.2)	5.9
Platelet inhibitors	11,037 (31.6)	1425 (31.0)	1.3	916 (29.7)	4.1	4117 (30.3)	1421 (30.9)	1.3	914 (29.7)	2.1
Lipid regulating drugs	5431 (15.5)	995 (21.6)	15.7	688 (22.3)	17.3	2681 (19.8)	993 (21.6)	4.6	687 (22.3)	5.1
Antipsychotics	4327 (12.4)	302 (6.6)	20	195 (6.3)	20.9	918 (6.8)	302 (6.6)	0.8	195 (6.3)	0.2
Antidepressants	3796 (10.9)	494 (10.7)	0.4	289 (9.4)	5	1400 (10.3)	492 (10.7)	1.3	288 (9.3)	4
Insulins	2598 (7.4)	300 (6.5)	3.6	205 (6.6)	3.1	1018 (7.5)	299 (6.5)	3.9	205 (6.7)	3.7
Antidiabetic drugs	7193 (20.6)	1066 (23.2)	6.2	733 (23.8)	7.6	3206 (23.6)	1065 (23.2)	1.1	733 (23.8)	1.3
Oral corticosteroids	1894 (5.4)	335 (7.3)	7.6	221 (7.2)	7.2	796 (5.9)	331 (7.2)	5.4	219 (7.1)	4.8
Nonsteroidal anti-inflammatory	3860 (11.0)	638 (13.9)	8.5	400 (13.0)	5.9	1981 (14.6)	633 (13.8)	2.4	396 (12.9)	5.8
drugs										
Surgical operation after hip	31,341 (89.7)	4451 (96.7)	28.2	2995 (97.1)	29.9	12,971 (95.6)	4443 (96.7)	5.8	2990 (97.0)	5.2
fracture			0 11					ſ	(001) LOC	L
Kesidency in nursing nome	(6.62) 2006	469 (10.2)	41.8	307 (9.9)	42.0	(7.7) 4501	469 (10.2)	0.3	307 (10.0)	Q
^a Absolute standardized difference com	npared with non	-exposed group	ć							

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Methods and the list of covariates used in the PS model is shown in Supporting Table 2. Each treated patient was matched with up to three nontreated patients using sequential greedy matching with a caliper of 0.2 standard deviation. Those who failed to match with a nontreated patient were excluded. To assess the quality of matching, absolute standardized differences (ASDs) in covariates between treatment groups were estimated. After matching, all covariates had an ASD <0.25 (Table 1), indicating that the covariates were well balanced.⁽²⁷⁾ Survival rate or disease-free survival rates of CVEs were plotted using the Kaplan-Meier method.⁽²⁸⁾ Hazard ratios (HRs) and 95% confidence intervals (Cls) were estimated using a conditional Cox proportional hazard model stratified by the matched pairs. To examine the association of treatment and risk of CVEs over time, follow-up for 1, 3, 5, and 10 years was reported.

Sensitivity analysis was conducted to investigate any residual and unmeasured confounding. First, the risk of CVE is expected to be the highest close to the time of hip fracture. Thus, patients with late treatment would have a much lower risk of CVE at the time of treatment, leading to bias toward any protective effect of treatment. We, therefore, performed a sensitivity analysis by excluding patients with late treatment, defined as the start of first treatment over 180 days from the time of discharge. The cutoff 180 days was used because we observed that mortality of hip fracture stabilized after 180 days (Fig. 1). In addition, patients with a short exposure of the drug would likely not have the beneficial effect of treatment. Therefore, another sensitivity analysis that excluded patients with treatment duration less than 30 days was conducted. For some very frail hip fracture patients, pharmacologic treatment may be perceived as nonbeneficial. Such practice may result in treatment of the subpopulation thought to have better survival prospects and long-term benefits in physical functioning. Thus, we performed a validation study to evaluate if those patients receiving treatment would have better survival prospects and physical functioning. Details are provided in the Supporting Methods.

Subgroup analysis was performed to evaluate the risk of outcomes by gender, history of cardiovascular disease (CVD; ICD-9: 390–495 Diseases of the circulatory system), history of diabetes, and type of surgical procedure for hip fracture.

R (R Foundation for Statistical Computing, Vienna, Austria; https://www.r-project.org/) and SAS (version 9.3; SAS institute, Inc., Cary, NC, USA) were used for all statistical analyses. A two-sided p value <0.05 was considered significant.

Results

Baseline characteristics

Between January 1, 2005 and December 31, 2013, 46,253 patients aged \geq 50 years were admitted via the emergency room with a new diagnosis of hip fracture. The top three causes of mortality in the first year following hip fracture are shown in Fig. 1. The risk of cardiovascular mortality was the highest in the first month after hip fracture (22.2%), and dropped to 11.6% after 1 year. Nonetheless, risk of pneumonia mortality and cancer mortality increased continuously.

Among 34,991 patients included in the final cohort (Fig. 2), 2868 (8.2%) were prescribed antiosteoporosis medication in the first year, and 4602 (13.2%, treated group) by study end date. The mean \pm SD age of the cohort was 82 \pm 9.3 years and 24,337 patients (69.6%) were female. After PS-matching, 4594 patients in the treatment group were matched with 13,568 non-exposed patients. The median (interquartile range [IQR]) follow-up times in non-exposed and treated groups were 1076 (1349) days and 1446 (1371) days, respectively. Among the treated patients, 3081 patients (67.1%) were exposed to alendronate (98% of patients treated weekly and 2% treated daily). Of these alendronate-treated patients, over 60% were prescribed the drug in the first year with a mean \pm SD starting time of 100 \pm 93.5 days after hip fracture. Compared to those patients who were not on treatment, the patients receiving treatment was not associated



Fig. 1. Trend in top three cause of death after hip fracture. Cardiovascular disease is defined as ICD-10 codes 100–109, 111, 113, 120–151.



Fig. 2. CONSORT flow diagram of the study.

with better physical functioning or survival prospect, with an odds ratio of 1.01 (95% Cl, 0.41 to 2.50); p = 0.98. Survival curves of alendronate treatment and risk of CVEs are presented in Fig. 3A–C.

Alendronate and risk of CVE

At 1-year follow-up, the incidence of cardiovascular mortality was 108.9 and 34.7 per 10,000 patient-years for the nontreated and alendronate-treated groups, respectively (Table 2). Alendronate was associated with a reduced risk of 1-year cardiovascular mortality (HR 0.33; 95% Cl, 0.17 to 0.65; p = 0.001) and incident MI (HR 0.55; 95% Cl, 0.33 to 0.89; p = 0.014). For incident stroke, a marginally significant reduction in risk was observed at 5 years follow-up (HR 0.82; 95% Cl, 0.67 to 1.00; p = 0.049) and 10 years follow-up (HR 0.83; 95% Cl, 0.69 to 1.01; p = 0.065) (Table

2). The protective association of alendronate and CVEs declined over time but remained statistically significant. Sensitivity analyses (Supporting Tables 3 and 4) revealed similar findings. In a subgroup analysis of only, women, BP use was associated with a reduced risk of cardiovascular mortality and incident MI, in a trend similar to the primary analysis, whereas in men there was association with cardiovascular mortality but not with incident MI. For incident stroke, association was not observed in both sexes (Table 3).

Other antiosteoporosis medications and risk of CVE

In a secondary analysis, similar but statistically more significant findings were observed for all N-BPs exposures (Table 2). Salcatonin had no association with CVEs at 1-year follow-up (Table 4) but significant increased risks of incident MI at 5 years



Fig. 3. Survival curves of the association of alendronate and risk of CVE. (A) Cardiovascular mortality. (B) Incident MI. (C) Incident stroke.

Table 2. Risk of CVE with N-BPs

Group	Subjects (n)	Events (<i>n</i>)	Mortality/incidence rate, per 10,000 person-years	Hazard ratio (95% Cl)	p
1-Year follow-up					
Cardiovascular m	nortality				
Non-exposed	13,568	130	108.9 (91–129.3)	1	-
Alendronate	3081	10	34.7 (16.6–63.7)	0.33 (0.17–0.65)	0.001
All N-BPs	3778	13	37 (19.7–63.2)	0.35 (0.20-0.63)	<0.001
Incident myocard	dial infarction				
Non-exposed	12,708	151	135.3 (114.5–158.6)	1	-
Alendronate	2998	20	71.4 (43.6–110.3)	0.55 (0.34–0.89)	0.014
All N-BPs	3679	22	64.4 (40.3–97.4)	0.51 (0.32–0.81)	0.004
Incident stroke					
Non-exposed	10,188	229	257.1 (224.9–292.7)	1	-
Alendronate	2696	49	194.6 (144–257.3)	0.78 (0.56–1.08)	0.133
All N-BPs	3299	56	182.8 (138.1–237.4)	0.70 (0.52–0.95)	0.022
3-Year follow-up					
Cardiovascular m	nortality				
Non-exposed	13,568	301	102.7 (91.4–115)	1	-
Alendronate	3081	36	47.3 (33.1–65.4)	0.48 (0.33–0.69)	<0.001
All N-BPs	3778	45	48.4 (35.3–64.8)	0.47 (0.34–0.66)	<0.001
Incident myocard	dial infarction				
Non-exposed	12,708	364	132.9 (119.6–147.3)	1	-
Alendronate	2998	57	77.2 (58.5–100.1)	0.63 (0.47–0.85)	0.002
All N-BPs	3679	64	71.1 (54.7–90.8)	0.58 (0.44–0.76)	<0.001
Incident stroke					
Non-exposed	10,188	526	242.5 (222.2–264.1)	1	-
Alendronate	2696	132	199.7 (167.1–236.9)	0.88 (0.71–1.09)	0.226
All N-BPs	3299	152	189.3 (160.4–221.9)	0.80 (0.66–0.98)	0.027
5-Year follow-up					
Cardiovascular m	nortality				
Non-exposed	13,568	386	99 (89.4–109.4)	1	-
Alendronate	3081	56	53.3 (40.3–69.2)	0.55 (0.40–0.75)	< 0.001
All N-BPs	3778	69	53.5 (41.6–67.7)	0.54 (0.41–0.72)	<0.001
Incident myocard	dial infarction				
Non-exposed	12,708	506	139 (127.2–151.7)	1	_
Alendronate	2998	100	98.4 (80–119.6)	0.70 (0.55–0.90)	0.005
All N-BPs	3679	121	96.9 (80.4–115.8)	0.69 (0.55–0.86)	0.001
Incident stroke					
Non-exposed	10,188	647	225.3 (208.3–243.4)	1	-
Alendronate	2696	168	185.1 (158.2–215.3)	0.82 (0.67–1.00)	0.049
All N-BPs	3299	198	1/8.5 (154.5–205.2)	0.77 (0.65–0.93)	0.006
10-Year follow-up					
Cardiovascular m	nortality				
Non-exposed	13,568	429	96.3 (87.4–105.8)	1	-
Alendronate	3081	78	63.9 (50.5–79.7)	0.59 (0.44–0.79)	< 0.001
All N-BPs	3//8	92	60.8 (49–74.5)	0.58 (0.44–0.75)	<0.001
Incident myocard	dial infarction	500		-	
Non-exposed	12,708	580	139.6 (128.5–151.5)		-
Alendronate	2998	123	104.2 (86.6–124.3)	0.71 (0.56–0.89)	0.004
All N-BPs	3679	145	99.1 (83.6–116.6)	0.67 (0.54–0.83)	<0.001
Incident stroke	10.100	<i></i>		-	
Non-exposed	10,188	694	212 (196.5–228.4)	1	-
Alendronate	2696	183	173.9 (149.6–200.9)	0.83 (0.69–1.01)	0.065
AII N-BPs	3299	220	169.6 (147.9–193.5)	0.79 (0.66–0.94)	0.008

N-BPs included alendronate, ibandronate, risedronate, and zoledronate. Patients with history of myocardial infarction or stroke were excluded from the analysis.

follow-up (HR 2.0; 95% CI, 1.16 to 3.46; p = 0.013) and at 10 years follow-up were observed (HR 2.0; 95% CI, 1.17 to 3.41; p = 0.011). Strontium ranelate showed no association with CVEs. Sensitivity analysis revealed similar findings (Supporting Tables 3 and 4).

Discussion

This is the first population-based study using a large electronic clinical patient record database to examine the risk of major CVE

and instant disconference		5	5	;								
		1-Year follow-up			3-Year follow-up			5-Year follow-up			10-Year follow-up	
Group <i>(n</i>)	Events (<i>n</i>)	Hazard ratio (95% CI)	d	Events (<i>n</i>)	Hazard ratio (95% CI)	ф	Events (<i>n</i>)	Hazard ratio (95% CI)	d	Events (<i>n</i>)	Hazard ratio (95% CI)	р
Cardiovascular mortality												
	v	0.27 (0.11–0.63)	0 003	7 7	0 44 (0 28-0 69)	~0.001	40	0 49 (0 33-0 72)	~0.001	50	053 (036-076)	/0007
Malo (n - 406)) ((20.0-11.0) (2.0		2	0.20 0.10 0.00	0.02	p r	(7/0-00) (10) (10) (10) (10) (10) (10) (10) (0.051		(0.10 00.0) 00.0	
History of any CVE	٩		0000	þ		0.0				n		10.0
No $(n = 984)$	m	1.51 (0.29–7.77)	0.622	5	0.42 (0.15–1.15)	0.092	8	0.41 (0.17–0.98)	0.045	14	0.41 (0.18–0.92)	0.03
Yes $(n = 1657)$	9	0.22 (0.09–0.56)	0.002	29	0.51 (0.31–0.83)	0.007	44	0.59 (0.37–0.92)	0.019	55	0.58 (0.38-0.9)	0.016
History of diabetes												
No (<i>n</i> = 2313)	7	0.40 (0.18–0.90)	0.028	22	0.42 (0.25-0.68)	<0.001	33	0.47 (0.31–0.72)	<0.001	47	0.52 (0.35-0.78)	0.001
Yes $(n = 386)$	-	0.25 (0.03–2.22)	0.215	7	0.78 (0.30–2.03)	0.603	11	0.89 (0.37–2.14)	0.796	15	0.81 (0.34–1.92)	0.639
Surgical operation												
Hip fixation ($n = 1770$)	S	0.24 (0.09-0.62)	0.003	20	0.44 (0.26–0.73)	0.002	34	0.57 (0.37–0.88)	0.011	45	0.6 (0.39–0.91)	0.018
Hip replacement	4	1.19 (0.23–6.08)	0.835	10	0.61 (0.26–1.42)	0.249	13	0.47 (0.21–1.07)	0.071	20	0.58 (0.28–1.2)	0.145
(n = 791)												
Incident myocardial infarction	on											
Sex												
Female ($n = 2344$)	10	0.41 (0.21–0.81)	0.011	35	0.46 (0.31-0.69)	<0.001	69	0.55 (0.4–0.75)	<0.001	86	0.56 (0.41-0.75)	<0.001
Male (<i>n</i> = 379)	8	1.09 (0.40–2.93)	0.865	14	1.08 (0.50-2.32)	0.845	19	1.02 (0.52–2.02)	0.954	20	1.02 (0.52-2.02)	0.954
History of any CVE												
No (<i>n</i> = 984)	4	0.58 (0.18–1.88)	0.363	∞	0.41 (0.18–0.91)	0.028	18	0.48 (0.26–0.89)	0.02	24	0.52 (0.29–0.94)	0.031
Yes (<i>n</i> = 1552)	12	0.53 (0.27–1.01)	0.054	42	0.64 (0.44–0.95)	0.025	70	0.7 (0.49–0.99)	0.044	82	0.68 (0.48–0.97)	0.031
History of diabetes												
No (<i>n</i> = 2255)	12	0.54 (0.28–1.03)	0.06	31	0.53 (0.35–0.81)	0.003	60	0.59 (0.42–0.83)	0.002	76	0.6 (0.43–0.83)	0.002
Yes $(n = 353)$	m	1.12 (0.22–5.62)	0.891	10	0.81 (0.32–2.02)	0.644	16	0.95 (0.41–2.23)	0.914	20	0.87 (0.38–2.01)	0.752
Surgical operation												
Hip fixation	13	0.65 (0.35–1.24)	0.19	29	0.69 (0.45–1.06)	0.09	52	0.77 (0.53–1.11)	0.163	62	0.8 (0.56–1.14)	0.212
(n = 1711)												
Hip replacement	S	0.74 (0.21–2.6)	0.64	19	0.93 (0.50–1.72)	0.814	30	0.92 (0.53–1.57)	0.748	35	0.93 (0.55–1.57)	0.791
(n = 752)												
Incident stroke												
Sex												
Female (<i>n</i> = 2106)	37	0.78 (0.53-1.15)	0.207	106	0.94 (0.73–1.21)	0.635	130	0.85 (0.67–1.07)	0.164	142	0.86 (0.69–1.08)	0.2
Male ($n = 301$)	7	1.18 (0.42–3.31)	0.757	17	0.77 (0.37–1.58)	0.475	22	0.76 (0.38–1.51)	0.425	23	0.72 (0.36–1.42)	0.341
History of any CVE												
No (<i>n</i> = 984)	6	0.77 (0.32–1.82)	0.547	32	1.11 (0.66–1.85)	0.693	39	1.01 (0.63–1.62)	0.968	42	0.91 (0.58–1.44)	0.701
Yes (<i>n</i> = 1180)	32	0.80 (0.50-1.26)	0.327	79	0.92 (0.67–1.26)	0.604	103	0.83 (0.62–1.11)	0.202	113	0.88 (0.66–1.17)	0.381
History of diabetes												
No (<i>n</i> = 2018)	31	0.82 (0.53–1.27)	0.371	85	0.85 (0.65–1.12)	0.259	112	0.79 (0.61–1.01)	0.065	123	0.79 (0.62–1.02)	0.068
Yes $(n = 250)$	S	0.39 (0.10–1.44)	0.158	17	0.87 (0.42–1.81)	0.712	22	0.9 (0.46–1.79)	0.772	25	0.96 (0.49–1.88)	0.909
Surgical operation												

Table 3. Subgroup Analysis Between Alendronate Treatment and CVE

lable 3. (Continued)												
		1-Year follow-up			3-Year follow-up			5-Year follow-up			10-Year follow-up	
Group (n)	Events (<i>n</i>)	Hazard ratio (95% CI)	٩	Events (<i>n</i>)	Hazard ratio (95% CI)	đ	Events (<i>n</i>)	Hazard ratio (95% CI)	đ	Events (<i>n</i>)	Hazard ratio (95% CI)	d
Hip fixation (<i>n</i> = 1467)	27	0.87 (0.55–1.40)	0.57	64	0.76 (0.55–1.05)	0.092	81	0.67 (0.5–0.9)	0.009	06	0.68 (0.51–0.92)	0.011
Hip replacement $(n = 644)$	σ	0.74 (0.31–1.77)	0.498	30	0.93 (0.54–1.61)	0.799	41	0.95 (0.58–1.57)	0.849	45	0.98 (0.6–1.61)	0.95
Patients with history of my	ocardial infa	rction or stroke were ex	cluded fro	om the and	alysis.							

in hip fracture patients with and without alendronate treatment. Patients prescribed alendronate treatment versus non-treatment had a significantly reduced risk of CVE. The association could endure for 10 years after fracture and was robust in various sensitivity analyses. Nonetheless, it appeared that the protective association was not evident for other classes of antiosteoporosis treatment.

Hip fracture is often under-treated with antiosteoporosis medication, which is the worldwide experience. In the current study, only 13.2% of patients received antiosteoporosis medication following hip fracture. Notably, we showed that post-hip fracture use of alendronate reduced cardiovascular mortality. This highlights the importance of initiating alendronate treatment after hip fracture.

Our findings were contrary to some studies where antiosteoporosis treatment was associated with an increased CVE risk, one of the reasons for the "Crisis in the Treatment of Osteoporosis."⁽²⁹⁾ Treatment with bisphosphonates has been associated with an increased risk of MI in patients with a history of fracture.⁽³⁰⁾ The population in that study was mainly male veterans (>95%), which could explain the discrepancy with our results. Indeed, the current study showed increased risk of incident MI in men but the association was not significant probably because of the relatively small numbers of men in the cohort.

A recent meta-analysis of RCTs on treatment of bisphosphonates reported a decreased risk of cardiovascular mortality but the association was not significant (pooled risk ratio: 0.81; 95% CI, 0.64 to 1.02).⁽¹¹⁾ The magnitude of effect sizes that we observed differed from those reported in the RCTs. Overestimation of treatment effect in PS-based observational studies is commonly reported.^(31,32) One possible reason for the discrepancy is the difference in the study populations. In the meta-analysis, four out of 10 trials targeted patients with cancer, whereas our cohort excluded these patients. Nevertheless, although PS matching has minimized the confounding in observational studies, selection bias due to unmeasured factors may still exist. For example, we cannot evaluate drug adherence by the patients, which is a common limitation of healthcare database research. Therefore, bias due to a "healthy adherer effect"(33) cannot be ruled out. In RCTs, intention-to-treat analysis is a common approach to address bias due to participants being lost-to follow up. However, some reviews suggested that this approach might underestimate treatment effect, resulting in larger discrepancies between RCTs and observational studies.^(34,35) On the other hand, no association of treatment and incident MI was shown in the meta-analysis. The studies included in the meta-analysis had different follow up periods ranging from 1 to 15 years. However, the current study showed that the protective association of treatment and CVEs declined over time. Such findings suggest that the studies with long follow-up periods in the meta-analysis may dilute the association, leading to the discrepancy of findings between the meta-analysis and the current study.

One RCT of zolendronate showed reduction of mortality after hip fracture only after the first year,⁽³⁶⁾ which was contrary to our findings. In the current study, most of the patients were treated with alendronate whereas only small number of cases were treated with other N-BPs. We cannot rule out the possibility that the protective association of alendronate and other N-BPs on the risk of CVE are different,

Table 4. Risk of CVE With Salcatonin and Strontium Ranelate

Group	Subjects (n)	Events (<i>n</i>)	Mortality/incidence rate, per 10,000 person-years	Hazard ratio (95% Cl)	p
1-Year follow-up					
Cardiovascular mortality					
Non-exposed	13,568	130	108.9 (91–129.3)	1	_
Salcatonin	535	8	201 (86.8–396)	2.33 (0.89-6.10)	0.084
Strontium ranelate	167	3	206.4 (42.6-603.1)	0.69 (0.15-3.24)	0.635
Incident myocardial infa	rction				
Non-exposed	12,708	151	135.3 (114.5–158.6)	1	-
Salcatonin	509	9	239.4 (109.5–454.5)	1.27 (0.56–2.89)	0.562
Strontium ranelate	160	3	215.3 (44.4–629.3)	1.14 (0.28–4.66)	0.854
Incident stroke					
Non-exposed	10,188	229	257.1 (224.9–292.7)	1	-
Salcatonin	459	11	329 (164.2–588.6)	1.64 (0.75–3.59)	0.218
Strontium ranelate	143	4	324.9 (88.5–831.8)	1.07 (0.31–3.72)	0.915
3-Year follow-up					
Cardiovascular					
mortality					
Non-exposed	13,568	301	102.7 (91.4–115)	1	-
Salcatonin	535	15	167.6 (93.8–276.5)	1.61 (0.82–3.15)	0.165
Strontium ranelate	167	6	166.9 (61.2–363.2)	1.26 (0.44–3.63)	0.671
Incident myocardial infa	rction				
Non-exposed	12,708	364	132.9 (119.6–147.3)	1	-
Salcatonin	509	21	249.2 (154.3–381)	1.84 (1.02–3.31)	0.042
Strontium ranelate	160	5	145.7 (47.3–339.9)	0.91 (0.31–2.66)	0.857
Incident stroke					
Non-exposed	10,188	526	242.5 (222.2–264.1)	1	-
Salcatonin	459	21	282.3 (174.8-431.6)	1.40 (0.80–2.47)	0.241
Strontium ranelate	143	10	330.5 (158.5–607.9)	1.24 (0.54–2.82)	0.616
5-Year follow-up					
Cardiovascular mortality					
Non-exposed	13,568	386	99 (89.4–109.4)	1	-
Salcatonin	535	17	152.1 (88.6–243.5)	1.59 (0.83–3.03)	0.163
Strontium ranelate	167	7	141.1 (56.7–290.7)	0.98 (0.37–2.60)	0.967
Incident myocardial infa	rction				
Non-exposed	12,708	506	139 (127.2–151.7)	1	-
Salcatonin	509	26	247.4 (161.6–362.4)	2.00 (1.16–3.46)	0.013
Strontium ranelate	160	8	168.6 (72.8–332.3)	1.13 (0.46–2.79)	0.786
Incident stroke					
Non-exposed	10,188	647	225.3 (208.3–243.4)	1	-
Salcatonin	459	23	248.9 (157.8–373.4)	1.25 (0.73–2.13)	0.415
Strontium ranelate	143	14	340.3 (186.1–571)	1.24 (0.54–2.82)	0.616
10-Year follow-up					
Cardiovascular					
mortality					
Non-exposed	13,568	429	96.3 (87.4–105.8)	1	-
Salcatonin	535	18	142.5 (84.4–225.2)	1.65 (0.87–3.12)	0.123
Strontium ranelate	167	8	137.1 (59.2–270)	1.10 (0.43–2.78)	0.841
Incident myocardial infa	rction				
Non-exposed	12,708	580	139.6 (128.5–151.5)	1	-
Salcatonin	509	29	243.2 (162.9–349.2)	2.00 (1.17–3.41)	0.011
Strontium ranelate	160	10	178.7 (85.7–328.6)	1.11 (0.49–2.52)	0.805
Incident stroke					
Non-exposed	10,188	694	212 (196.5–228.4)	1	-
Salcatonin	459	25	241 (155.9–355.7)	1.29 (0.76–2.19)	0.344
Strontium ranelate	143	14	291.8 (159.5–489.6)	1.24 (0.54–2.82)	0.616

Patients with history of myocardial infarction or stroke were excluded from the analysis.

even though they belong to the same drug class. Given the relatively small number of zolendronate users and other N-BPs, we were not able to test this hypothesis. Further study with larger sample size of N-BPs is warranted.

Confounding by indication affects the validity of pharmacoepidemiology studies, thus we performed an analysis to evaluate if the cardiac-protective association was observed for two other antiosteoporosis medications with a different mechanism-of-action. These medications were associated with a nonsignificant or non-robust increased CVE risk (Table 4), especially for strontium ranelate that has been shown to be associated with increased CVE risk. Such a finding could be due to our limited sample size, or the null association, which was reported in two large population-wide studies.^(37,38) In addition, patients with short-term or late treatment may bias the effect of treatment because short exposure of the drug would have little beneficial effect on CVEs and patients with late treatment would have a much lower risk of CVE at the time of treatment. To address the bias, we excluded these patients in the sensitivity analysis and the results remained robust. Furthermore, we showed that the prescription of bisphosphonate did not differ according to the physical functioning of the patients, suggesting that the observed association with alendronate was not due to better patientcare, survival prospect, or physical functioning.

The association of alendronate and reduced risk of CVEs could be explained by the extra-mineral and skeletal effect of N-BPs. N-BPs target FPPS in the mevalonate pathway that belongs to the same pathway as statins. Thus N-BPs have a cholesterollowering effect.⁽³⁹⁾ Bisphosphonates can also modulate ion channels in cardiac myocytes,^(40,41) regulate and inhibit vessel pathogenesis,⁽⁴²⁾ and has an anti-inflammatory effect.⁽⁴³⁾ Animal studies have shown that N-BPs attenuate diastolic dysfunction following MI,⁽⁴⁴⁾ improve cardiac properties, and reduce severity of CVE.^(9,10)

The current study has important clinical implications. It is well established that there is a worldwide crisis in the treatment of osteoporosis,⁽²⁹⁾ due to patients' awareness of the potential side effects. This leads to under-use of the treatment in hip fracture patients, even though multiple clinical guidelines recommend the use. If our findings are further validated, optimal uptake of antiosteoporosis medication can be encouraged. In addition, our study has important implications for RCT design. RCT of new antiosteoporosis agents often use alendronate as a comparator; eg, the RCT of romosozumab.⁽⁴⁵⁾ The US Food and Drug Administration (FDA) has recently requested more data before reaching to decision on whether to approve the osteoporosis drug romosozumab, due to the excess cardiovascular adverse events in the romosozumab arm compared with the alendronate arm.⁽⁴⁶⁾ In light of these important deliberations, our results provide evidence that such differences in cardiovascular adverse events could be potentially related to protective association of alendronate, rather than an increase in cardiovascular adverse events related to romosozumab use.

Our study has several strengths. To the best of our knowledge, this is the first large contemporary analysis of real-world clinical practice to compare the risk of CVE among hip fracture patients with and without antiosteoporosis treatment, and complements results from RCTs. All records in the CDARS were validated with high accuracy,⁽⁴⁷⁾ and is a powerful platform from which to conduct large-scale, postmarketing, drug surveillance studies.^(16,17,48,49) This study was also carefully designed. Patients with previous exposure to

antiosteoporosis treatment were excluded to avoid a residual effect of treatment. Similarly, patients with history of hip fracture and cancer were excluded. Using a time-dependent PS matching method, the potential confounding factors between treated and untreated groups were minimized, and we included multiple sensitivity analyses to further reduce the confounding bias. Immortal bias due to delay of treatment was adjusted using the robust method.⁽²³⁾

There were several limitations in the current study. First, similar to other studies utilizing the healthcare record, over-thecounter products by a non-HA pharmacy are not captured by the CDARS. Nonetheless, patients with chronic diseases who require long-term treatment commonly use the service of HA because the medication cost is highly subsidized. Therefore, the impact of uncaptured medications should be minimal. Second, although we excluded patients with prescription records 2 years prior to hip fracture, the residual effect of antiosteoporosis medication may exceed this time, although the effect should be minimal. Third, the effect of bone mineral density (BMD) on CVE is unknown. One would expect that patients with a lower BMD would be likely to be prescribed antiosteoporosis treatment. Nonetheless, it is known that low BMD is associated with higher CVE risk. Therefore, even if BMD affects treatment decisions, it would have led to underestimation, not overestimation, of the treatment effect. Fourth, data were not available on body mass index, blood pressure, blood lipids, and smoking, which are risk factors for CVEs. To address the concern, we included the diagnoses of overweight and obesity, hypertensive diseases, and hyperlipidemia in the PS model as surrogate markers of these factors. However, residual bias is still possible. Similarly, other potential confounding factors are not captured in CDARS, such as emigration, vitamin D and calcium supplementation use. However, it is expected that these confounding factors may not confer large effects on the clinical outcomes, especially in a short period of time (eq, 1-year cardiovascular mortality, MI, and stroke).

In conclusion, osteoporosis is undertreated among hip fracture patients. The use of alendronate was associated with a reduced risk of cardiovascular mortality, MI, and stroke. If the results are further validated, the initiation of alendronate treatment in patients with hip fracture is encouraged.

Disclosures

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